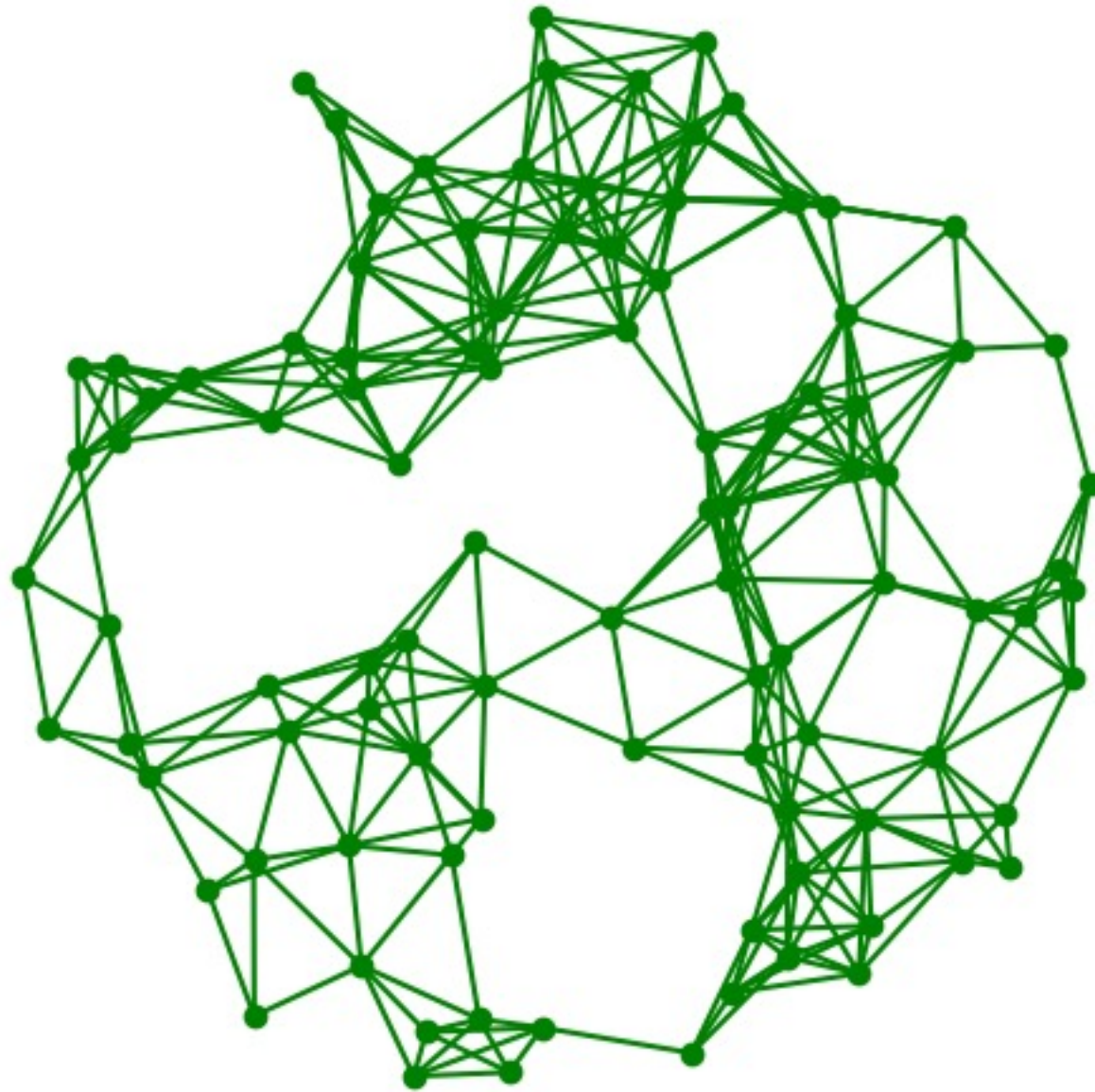


# Dynamics and inference for voter model processes



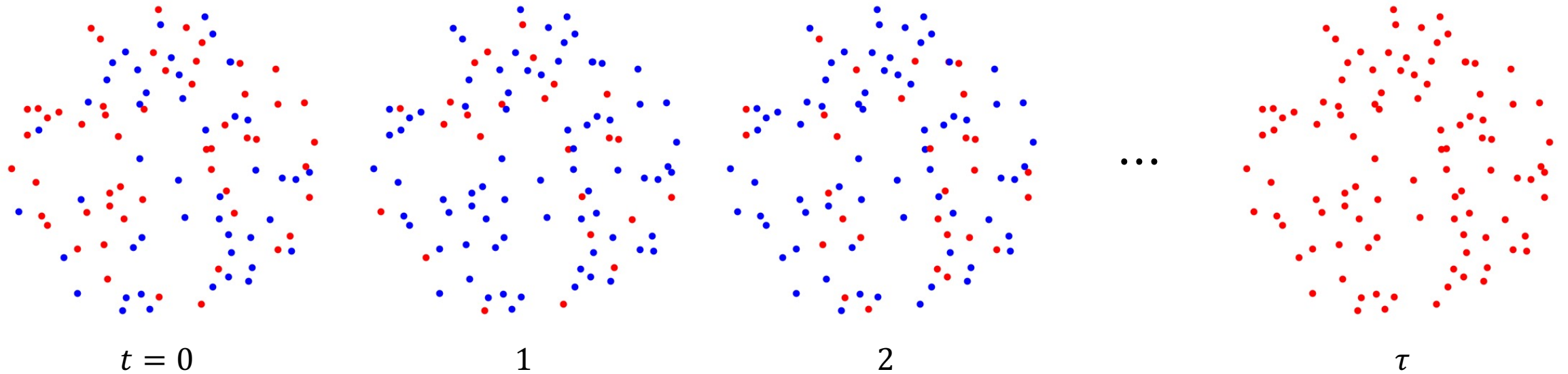
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Joint work with Kaifang Zhou

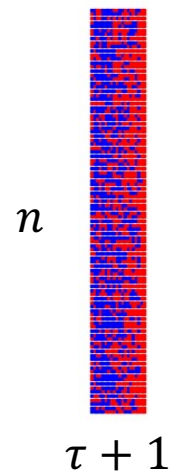


Nodes engage in pairwise interactions and upon interaction update their states  
Interactions restricted by a communication graph

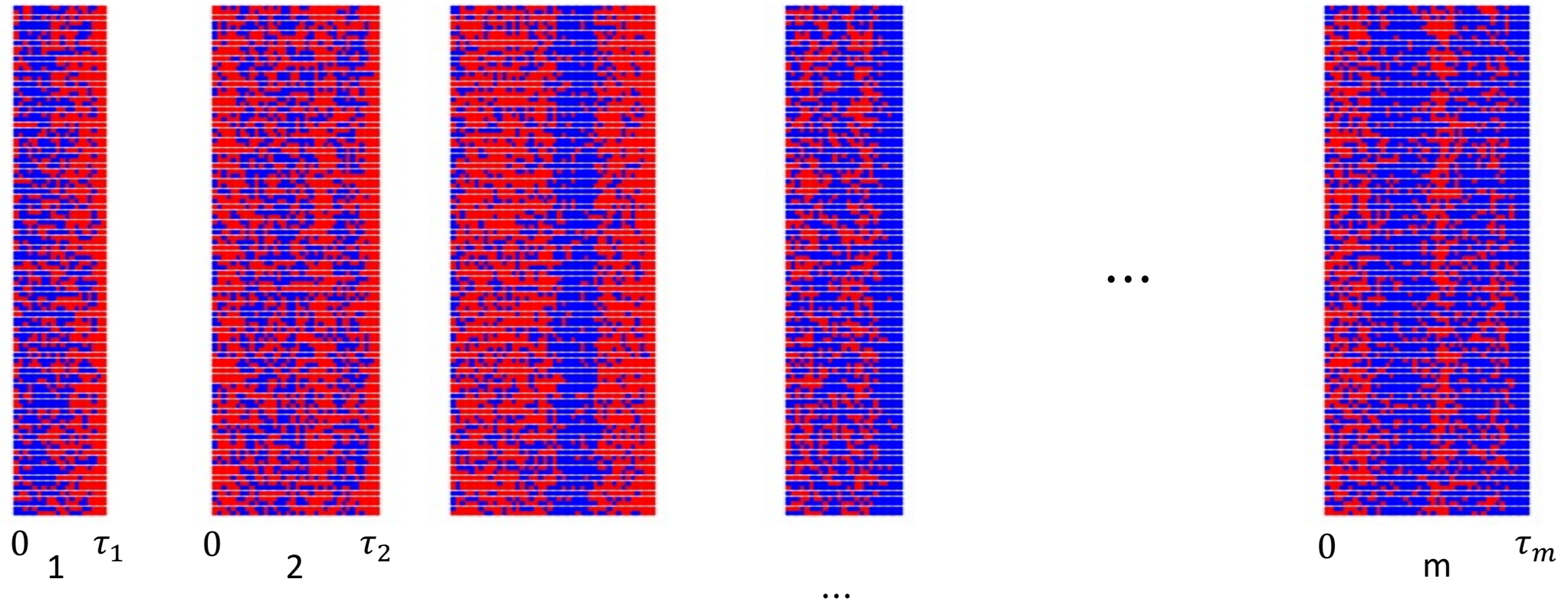
# Observations: node states



Observations:



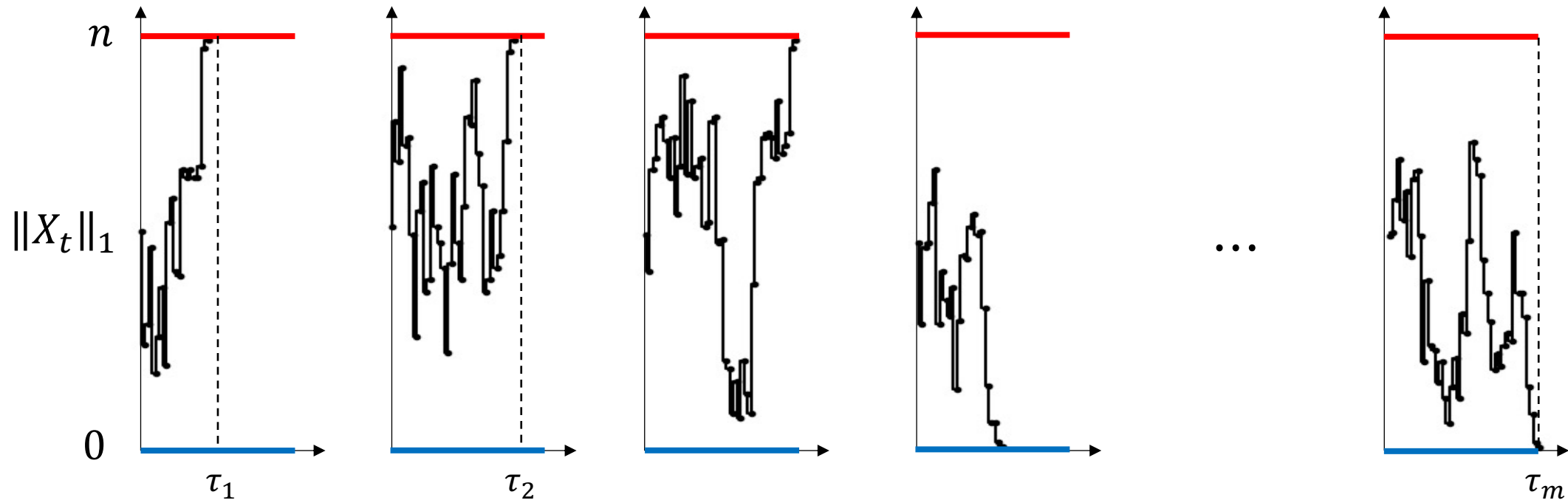
# Observations: node states (cont'd)



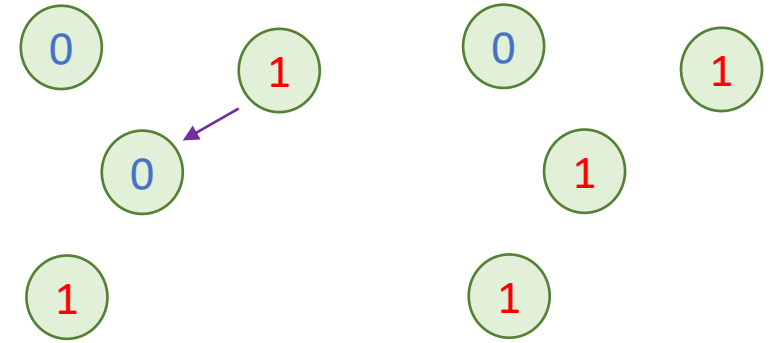




# Consensus times



# Voter model



- Each node state of value 0 or 1
- At an interaction time a node adopts the state of randomly sampled neighbor
- Classical model: Holley and Liggett (1975), Liggett (1985), ...
- Studied under different assumptions about interaction time instances
- Our focus:
  - Discrete-time model, in each time step every node updates its state
  - Random neighbor selection with probabilities  $A$  where the  $u$ -th row is the sampling probability distribution of node  $u$

# Dynamics and inference

- Much work has been devoted to studying dynamics of voter model processes
  - Hitting probabilities of absorption states (consensus), assuming convergence is to a consensus state
  - Hitting time (consensus time)
- Much less is known about parameter estimation (node sampling probabilities) from observed data



# Voter model process

- Initial state  $X_0 \sim \mu$ , and

$$X_{t+1,u} | X_t \sim \text{Ber}(a_u^\top X_t) \quad \text{for } t = 0, 1, \dots, \quad u \in \{1, \dots, n\}$$

where  $A = (a_1, \dots, a_n)^\top$  is the model parameter

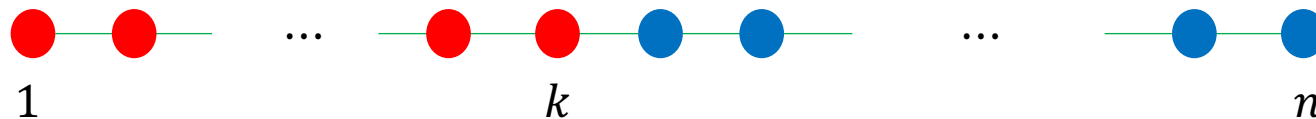
- Or, equivalently,

$$X_{t+1} = Z_{t+1} X_t \quad \text{for } t = 0, 1, \dots,$$

where  $Z_1, Z_2, \dots$  are i.i.d. random stochastic matrices in  $\{0, 1\}^{n \times n}$ ,  $\mathbf{E}[Z_1] = A$

# Parameter estimation is “hard”

- Path example
  - At each time step a random node initiates interaction
  - Communication graph is a path
  - Initial state:  $k$  nodes on one end of path in state  $1$ , other nodes in state  $0$



- An interaction is *informative* only if initiated by a node with disagreeing neighbors
- Expected number of informative interactions =  $k \left( \log \left( \frac{n}{k} \right) + \Theta(1) \right)$
- Number of unknown parameters:  $\Theta(n)$

# Challenges

- Number of observations is a priori random for any fixed number  $m$  of voter model process realizations
- Existing work focused on inference for stationary stochastic processes for a fixed number of observation points
- Some related work
  - High-dimensional generalized linear autoregressive models: Hall et al (2019)
  - Sparse multivariate Bernoulli processes in high dimensions: Pandit et al (2019)
  - Network vector autoregression: Zhu et al (2017)
  - Inferring graphs from cascades: Pouget-Abadie and Horel (2015)

# Limit to a consensus state

- **Thm** [Hassin and Peleg, 2001] For any  $A$  corresponding to adjacency of a nonbipartite graph, for any initial state  $x$ ,

$$\lim_{t \rightarrow \infty} \mathbf{P}[X_t = \mathbf{1}] = 1 - \lim_{t \rightarrow \infty} \mathbf{P}[X_t = \mathbf{0}] = \pi^\top x$$

where  $\pi$  is the stationary distribution for  $A$ , i.e.  $\pi^\top = \pi^\top A$

- Consensus states  $C = \{\mathbf{0}, \mathbf{1}\}$

# Consensus time

- Hassin and Peleg (2001):  $\mathbf{E}[\tau] = O(m(G) \log(n))$  where  $m(G)$  is the worst-case expected meeting time for two random walks on  $G$
- Berenbrink et al (2016):  $\mathbf{E}[\tau] = O\left(\frac{1}{\Phi(G)} \frac{d(V)}{d_{\min}}\right)$  for lazy random walk
- Kanade et al (2019):  $m(G) = O\left(\frac{1}{\Phi(G)} n d_{\max} \log(d_{\max})\right)$  for lazy random walks
- **Lazy random walk:** with probability  $\frac{1}{2}$  moves to a randomly chosen neighbor and otherwise remains at the current node

# Graph conductance

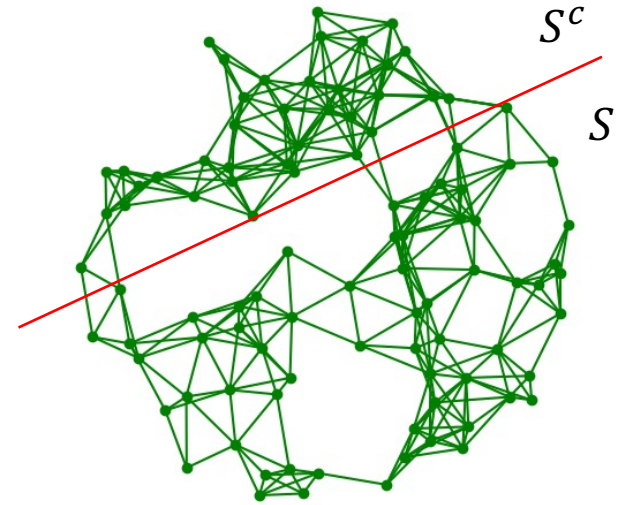
- Graph conductance of graph  $G = (V, E)$ ,

$$\Phi(G) = \min_{S \subset V: 0 < |S| < n} \frac{|E(S, S^c)|}{\min\{d(S), d(S^c)\}}$$

where  $E(S, S^c)$  is the set of edges with vertices in  $S$  and  $S^c$ ,  $d(S)$  is the sum of degrees of nodes in  $S$

- Cheeger's inequality:  $\frac{\lambda_2}{2} \leq \Phi(G) \leq \sqrt{2\lambda_2}$  where  $\lambda_2$  is the second smallest eigenvalue of the normalized Laplacian matrix

$$L = I - D^{-\frac{1}{2}}AD^{-\frac{1}{2}}$$





# Consensus time (cont'd)

- Cooper and Rivera (2016):

$$\mathbf{E}[\tau] \leq \frac{64}{\Psi_A}$$

where

$$\Psi_A = \pi^* \tilde{\Psi}_A$$

and

$$\tilde{\Psi}_A = \min_{x \in \{0,1\}^n \setminus C} \frac{E[|\sum_{u=1}^n \pi_u (x_u - \sum_{v=1}^n Z_{u,v} x_v)|]}{\min\{\pi^\top x, 1 - \pi^\top x\}}$$

# Expected consensus time bound

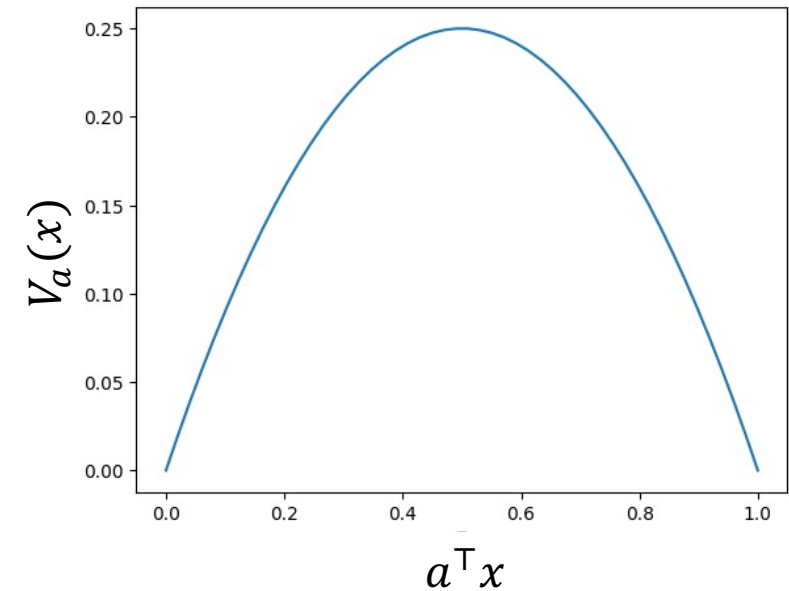
- **Thm** For every initial state  $x \in \{0,1\}^n$ ,

$$\mathbf{E}_x^0[\tau] \leq \frac{1}{\Phi_A} \log \left( \frac{1}{2\pi^*} \right)$$

where

$$\Phi_A = \min \left\{ \frac{\sum_{u=1}^n \pi_u^2 V_{a_u}(x)}{\pi^\top x (1 - \pi^\top x)} : x \in \{0,1\}^n, x \notin C \right\}$$

and  $\pi^* = \min\{\pi_u : u = 1, \dots, n\}$



# Comments on $\Phi_A$

- Fact:  $0 < \Phi_A \leq 1$
- For  $A$  according to graph  $G$ , i.e.  $a_{u,v} = 1/d_u$  for  $(u, v) \in E$

$$\Phi_A = \min_{S \subset V: 0 < |S| < n} \frac{|E_2(S, S^c)|}{d(S)d(S^c)}$$

where  $E_2(S, S^c)$  is the set of paths of length equal to two edges, connecting a vertex in  $S$  and a vertex in  $S^c$

# Examples

- Complete graph  $K_n$ :

$$\Phi_A = \frac{n-2}{(n-1)^2} = \frac{1}{n} (1 + o(1))$$

- Cycle  $C_n$ :

$$\Phi_A = 4 \frac{1}{n^2} (1 + o(1))$$

# Relations between $\Phi_A$ , $\Phi(G)$ , $\Psi_A$

- Assume  $a_{u,v} = \frac{1}{2} \mathbf{1}_{\{u=v\}} + \frac{1}{2} \frac{1}{d_u} \mathbf{1}_{\{(u,v) \in E\}}$  (lazy random walk)

- Then,

$$\frac{1}{\Phi_A} \leq 2 \frac{d(V)}{d_{\min}} \frac{1}{\Phi(G)} \quad \text{and} \quad \frac{1}{\Phi(G)} \leq \frac{1}{\tilde{\Psi}_A}$$

- Hence,

$$\frac{1}{\Phi_A} \leq 2 \frac{1}{\Psi_A}$$

# Exponential moment bound

- **Thm** For any  $x \in \{0,1\}^n$  such that  $x \notin C$  and any  $\theta \in \mathbf{R}$  such that  $(1 - \Phi_A)e^\theta \leq 1$ , we have

$$\mathbf{E}_x^0[e^{\theta\tau}] \leq \frac{V_\pi(x)}{\min_{z \in \{0,1\}^n \setminus C} V_\pi(z)}$$

Proof: Follows from a general result for Markov chain hitting times.

Let  $\tau_S = \min\{t > 0: X_t \in S\}$

Assume that  $V: \mathcal{X} \rightarrow [1, \infty)$  is a measurable function that satisfies, for some set  $C$  and  $\lambda < 1$ ,  $\mathbf{E}[V(X_1) \mid X_0 = x] \leq \lambda V(x)$  for all  $x \notin C$ . Then,  $\mathbf{E}_x[\lambda^{-\tau_C}] \leq V(x)$ .



# A probability bound

- **Thm** Let  $\tau_1, \dots, \tau_m$  be consensus times of  $m$  independent realizations of voter model processes with parameter  $A$  with independent initial states according to arbitrary distributions. Then, for any  $a \geq 0$ ,

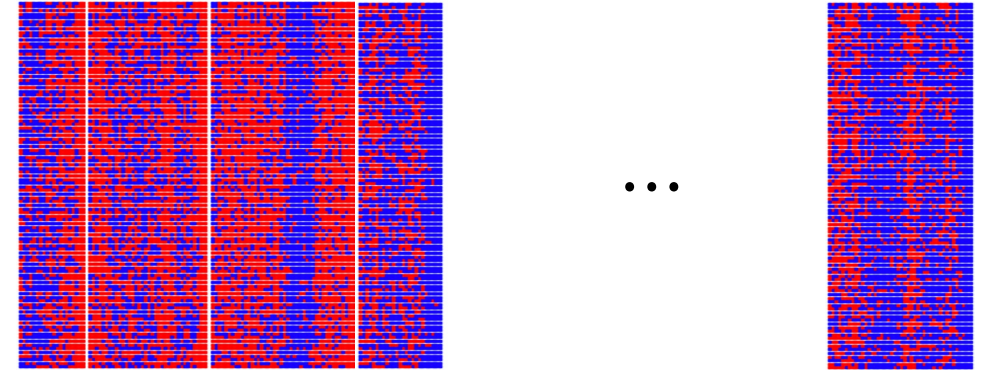
$$\mathbf{P}^0 \left[ \sum_{i=1}^m \tau_i \geq ma \right] \leq \left( \frac{\mathbf{E}^0[V_\pi(X_0)]}{\min_{z \in \{0,1\}^n \setminus C} V_\pi(z)} (1 - \Phi_A)^a \right)^m$$

It follows that for any  $\delta \in (0,1]$ , with probability at least  $1 - \delta$ ,

$$\sum_{i=1}^m \tau_i \leq \frac{1}{\Phi_A} \left( m \log \left( \frac{1}{2\pi^*} \right) + \log \left( \frac{1}{\delta} \right) \right)$$

# Parameter estimation

- Data:  $X = \left( X_0^{(1)}, \dots, X_{\tau_1}^{(1)}, \dots, X_0^{(m)}, \dots, X_{\tau_m}^{(m)} \right)^\top$



- We consider maximum likelihood estimation:

$$\hat{A} \in \arg \min_{A \in \Theta} \{ \mathcal{L}(A; X) \}$$

$$\text{where } \mathcal{L}(A; X) = \underbrace{-\ell(A; X)}_{\text{negative log-likelihood function}} + \underbrace{\lambda_m \|A\|_{1,1}}_{\text{regularization}}$$

negative log-likelihood function      regularization

$$\|B\|_{p,q} = \left( \|b_1\|_q^p, \dots, \|b_n\|_q^p \right)^{1/p}$$

# Parameter estimation bound

- **Thm** Consider the voter model process with parameter  $A^*$  with support size  $s$  and  $a_{u,v}^* \geq \alpha$  whenever  $a_{u,v}^* > 0$  for some  $\alpha > 0$ . Assume that  $\hat{A}$  is a minimizer of  $\mathcal{L}(A; X)$  with the regularization parameter

$$\lambda_m = 2\sqrt{2} \frac{c_{n,\pi^*}}{\alpha\sqrt{\Phi_{A^*}}} \sqrt{m}$$

and  $m$  is sufficiently large (precise condition omitted). Then, for some constant  $c > 0$ , with probability at least  $1 - 5/n$ ,

$$\|\hat{A} - A^*\|_F^2 \leq c \frac{sc_{n,\pi^*}^2}{\alpha^2 (\Phi_{A^*} \mathbf{E}^0[\tau])^2 \lambda_{\min}(\mathbf{E}[X_0 X_0^\top])^2} \Phi_{A^*} \frac{1}{m}$$

where

$$c_{n,\pi^*}^2 = \left( \log\left(\frac{1}{2\pi^*}\right) + \log(2n^3) \right) \log(4n^3)$$

# Proof sketch

- Proof is based on the framework of M-estimators with decomposable regularizers (Negahban et al 2012, Wainwright 2019)

- **Thm** Assume that loss function  $\mathcal{L}(A; X)$  has the regularization parameter such that

$$(C1) \quad \lambda_m \geq 2 \|\nabla \ell(A^*)\|_\infty$$

and

(C2) for some  $S \subseteq V^2$ ,  $-\ell(A; X)$  satisfies the restricted strong convexity (RSC) condition relative to  $A^*$  and  $S$  with curvature  $\kappa > 0$  and tolerance  $\gamma^2$

Then,

$$\|\hat{A} - A^*\|_F^2 \leq 9|S| \left(\frac{\lambda_m}{\kappa}\right)^2 + \left(2\gamma^2 \frac{1}{m} + 4\|A_{S^c}^*\|_{1,1}\right) \frac{\lambda_m}{\kappa}$$

# RSC condition

- A loss function  $\mathcal{L}$  is said to satisfy the RSC relative to  $A^*$  and  $S$  with curvature  $\kappa > 0$  and tolerance  $\gamma^2$  if

$$\mathcal{E}(\Delta) \geq \kappa \|\Delta\|_F^2 - \gamma^2 \text{ for all } \Delta \in \mathcal{C}(S; A^*)$$

where

$$\mathcal{E}(\Delta) = \mathcal{L}(A^* + \Delta) - \mathcal{L}(A^*) - \nabla \mathcal{L}(A^*)^\top \text{vec}(\Delta) \quad (\text{first-order Taylor error})$$

and

$$\mathcal{C}(S; A^*) = \left\{ \Delta: \|\Delta_{S^c}\|_{1,1} \leq 3 \|\Delta_S\|_{1,1} + 4 \|A_{S^c}^*\|_{1,1} \right\}$$

# Condition (C1)

- **Lem** For any  $\delta \in (0,1]$  and any  $m \geq 1$  independent realizations of the voter model process with parameter  $A^*$  and initial value distribution  $\mu$ , with probability at least  $1 - \delta$ ,

$$\|\nabla \ell(A^*)\|_\infty \leq \sqrt{2} \frac{1}{\alpha} \frac{1}{\sqrt{\Phi_{A^*}}} \sqrt{m} c_{n,\delta,\pi^*}(m)$$

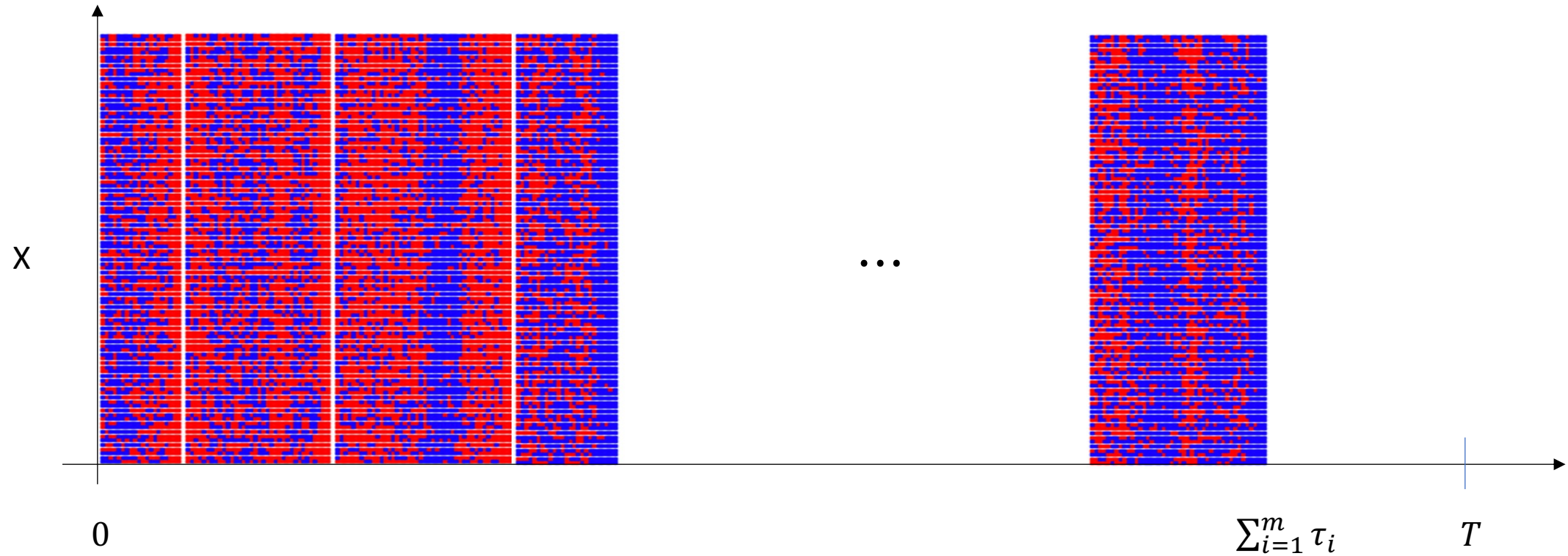
where

$$c_{n,\delta,\pi^*}(m)^2 = \left( \log\left(\frac{1}{2\pi^*}\right) + \frac{1}{m} \log\left(\frac{2n^2}{\delta}\right) \right) \log\left(\frac{4n^2}{\delta}\right)$$

Proof: Using a truncation argument, consensus time probability tail bound, and Azuma-Hoeffding's inequality for bounded-difference martingale sequences



# Truncation argument in a picture



$$\mathbf{P} \left[ E_{\sum_{i=1}^m \tau_i} \right] \leq \mathbf{P}[E_T] + \mathbf{P}[\sum_{i=1}^m \tau_i \geq T]$$



Some event

# Condition (C2)

- Show

$$\mathcal{E}(\Delta) \geq h(\Delta; X) := \sum_{i=1}^m \sum_{t=0}^{\tau_i-1} \sum_{u=1}^n \left( \Delta_u^\top X_t^{(i)} \right)^2$$

- Then show **(C2')**:  $h(\Delta; X)$  satisfies the RSC condition with high probability

# Condition (C2')

- **Step 1:**  $\mathbf{E}^0[h(\Delta; X)] \geq \kappa_1 \|\Delta\|_F^2$  for all  $\Delta$  where

$$\kappa_1 \leq m \mathbf{E}^0[\tau] \lambda_{\min}(\mathbf{E}[X_0 X_0^\top])$$

- **Step 2:** For any  $\delta \in (0, 1/2]$ , any  $S$  such that  $|S| \leq s$  and any  $\Delta \in \mathcal{C}(S; A^*)$ ,  $h(\Delta; X) \geq \frac{\kappa_1}{2} \|\Delta\|_F^2$  with probability at least  $1 - \delta$  provided that

$$m \geq \frac{s^2}{\Phi_{A^*} \mathbf{E}^0[\tau]^2 \lambda_{\min}(\mathbf{E}[X_0 X_0^\top])^2} c_{\delta, \pi^*}(m)$$

where

$$c_{\delta, \pi^*}(m) = 8 \left( \log \left( \frac{1}{2\pi^*} \right) + \frac{1}{m} \log \left( \frac{2}{\delta} \right) \right) \log \left( \frac{2}{\delta} \right)$$

# Condition (C2') cont'd

- **Step 3:** Show that

$$\mathbf{P}[h(\Delta; X) \geq \kappa' \|\Delta\|_F^2 - \gamma'^2 \text{ for all } \Delta \in \mathcal{C}(S; A^*)] \geq 1 - \frac{4}{n}$$

To show this, we apply some set covering arguments and combine with the bound in Step 2

# Conclusion

- Shown a new bound on consensus time, in expectation and probability
- Shown new parameter estimation bounds for absorbing voter model processes, obtained by leveraging the consensus time bounds